# What Do Consumers Consider Before They Choose?

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#### Introduction

- When estimating consumer demand models we usually assume that consumers consider all the alternatives that we as the researcher see
- Lots of evidence that the assumption of full consideration is violated in reality for many applications of interest
- We should care about this for a number of reasons, e.g:
  - Cannot predict the impact/evaluate the benefits of making consumers aware of a wider set of alternatives
  - Biased estimates of preference parameters with implications for welfare analysis

#### Introduction

- An exception: the literature on "consideration sets" consumers might only consider an (unobserved) subset of alternatives
- Popular in marketing and a growing applied literature for providing a "simple" way to introduce unobserved choice sets
  - ► (Behavioural) decision theory provides a rich set of models: see, e.g. Masatlioglu et al (2012) and Cattaneo et al (2018)
  - ▶ Default specific: Ho, Hogan & Scott-Morton (2016); Heiss, McFadden, Winter, Wuppermann & Zhou (2016); Moshkin & Shachar (2002)
  - ► Alternative specific: Goeree (2008); Manzini and Mariotti (2012); Conlin and Mortimer (2013); Honka et al (2015); Gaynor, Propper & Seiler (2016)

## Challenge

- Wider application of these models has been held back by the difficulty of separately identifying "utility" and "consideration probability" parameters from observational data
- Two main strategies pursued to date:
  - Auxiliary data: can we collect additional data on what options consumers considered?
  - 2. Exclusion restrictions: are there exogenous variables excluded from utility and from process generating consideration?

## This Paper

- In this paper we show that the restrictions from economic theory are sufficient for identification in many applied settings of interest
- Our approach relies on exploiting asymmetries in the "Slutsky" matrix
  - Changes in the characteristics of products impact the probability that you consider a good and not just utility
  - ► There is a particular pattern of cross-price asymmetries and violations of nominal illusion that are characteristic of a lack of consideration
  - Inspired by the theoretical work of Gabaix (2014) on inattention to characteristics although our focus is on inattention to goods

## This Paper

- ▶ Different strategy to that pursued in other current working papers on identification of consideration set models:
  - Crawford, Griffith & laria: results specific to Logit errors and rely on some assumptions about stability of choice sets over time
  - Dardanoni, Manzini, Mariotti & Tyson: limited allowance for preference heterogeneity
  - Cattaneo, Ma, Masatlioglu & Suleymanov: deterministic preferences but weaker assumptions on consideration
  - Barseghyan, Coughlin, Molinari & Teitelbaum: weaker assumptions on preference heterogeneity and consideration leading to set identification results

## This Paper

- Bring a parametric version of our framework to data to show that the variation at heart of our identification result is important for driving empirical results
  - Indirect inference estimator in which auxiliary model allows for cross derivative asymmetries
  - Structural parameters chosen to match the reduced form asymmetries
- ▶ Lab validation: can we recover the process generating consideration sets from choice data?
- Medicare Part D: to what extent is inertia driven by switching costs or lack of consideration?
  - Used to evaluate a proposed 'active default' policy

#### Outline

- I General Set-Up
- II Asymmetry-Based Identification
- **III** Estimation
- IV Experimental Validation
- V Field Application

## Basic Set-Up: Preferences

- Imagine that we are in a full-information environment
- ► Consumer *i* selects the good 0, ..., *J* that gives her the highest utility
- ▶ Utility is a function of a good's characteristics,  $\mathbf{x} \in \mathbb{R}^K$ , plus a random error

$$u_{ij} = v_j(\mathbf{x}_j) + \epsilon_{ij}$$
  
=  $\beta p_j + w_j(\mathbf{z}_j) + \epsilon_{ij}$ 

- Here assume quasi-linearity but show can be (partially) relaxed within main paper
- Proof extends naturally to allow for individual heterogeneity through a random coefficient

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## Basic Set-Up: Preferences

► The probability that a consumer chooses some good *j* is then:

$$\begin{array}{lcl} \textit{Pr}(\textit{i} \; \textit{chooses} \; \textit{j}) & = & \textit{Pr} \; \big( \textit{u}_{\textit{ij}} > \textit{u}_{\textit{ij'}} \quad \forall \textit{j'} \neq \textit{j} \big) \\ s^{\star}_{\textit{j}} & = & \textit{Pr} \; \big( \epsilon_{\textit{ij'}} < \textit{v}_{\textit{j}} + \epsilon_{\textit{ij}} - \textit{v}_{\textit{j'}} \quad \forall \textit{j'} \neq \textit{j} \big) \end{array}$$

**Example:** when  $\epsilon_{ij}$  is distributed Type 1 Extreme Value, we get the popular logit model

$$s_{j}^{\star} = \frac{exp(v_{j})}{\sum_{j'=1}^{J} exp(v_{j'})}$$

NB We allow for correlated unobservables in utility!

## Basic Set-Up: Consideration

- ► A consumer may not consider all goods in her choice set
- ► Good-0 represents an "inside" or "outside" default
- ▶ Let P(J) represent the power set of all goods, with any given element indexed by C
- ▶ Set of consideration sets containing good *j* is given as:

$$\mathbb{P}(j) = \{ C : C \in \mathcal{P}(J) \quad \& \quad j \in C \quad 0 \in C \}$$

## Basic Set-Up: Consideration

- Need some restrictions on consideration probabilities to achieve identification
- ► Two main classes of consideration set model found in the applied literature:
  - ▶ **Default specific**: with some probability  $\mu(\mathbf{x}_0)$  you consider the full choice set, otherwise you only consider a (known) default option
  - ► Alternative specific: you consider good j with probability  $\phi_j(\mathbf{x}_j)$
- We consider a general framework that subsumes both of these classes of model
- NB throughout this presentation will be assuming independence of unobservables driving utility and consideration

► In our model, observed choice probabilities take the form:

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▶ In our model, observed choice probabilities take the form:

$$\mathbf{s}_0 = (1 - \mu) + \mu \sum_{\mathbf{C} \in \mathbb{P}(0)} \prod_{l \in \mathbf{C}} \phi_l \prod_{l' \notin \mathbf{C}} (1 - \phi_{l'}) \, \mathbf{s}_0^{\star}(\mathbf{C})$$

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$$s_j = \mu \sum_{C \in \mathbb{P}(j)} \prod_{l \in C} \phi_l \prod_{l' \notin C} (1 - \phi_{l'}) s_j^*(C)$$

for j > 0

#### **Extensions**

- Dependence of  $\phi_j$  on the characteristics of the default product
- Independence of unobservables influencing utility and attention implicit in the background
  - Consider case of finite set of "types"
  - Require exclusion restrictions for identification but fewer than if ignored results in this paper
- Asymmetries and nominal illusion results that we will now develop imply imperfect consideration in wider class of models

#### Outline

- I General Set-Up
- **II Asymmetry-Based Identification**
- III Estimation
- IV Experimental Validation
- V Field Application

## Symmetry

 With full consideration, choice probabilities will satisfy a symmetry restriction

$$\frac{\partial s_j^{\star}}{\partial p_k} = \frac{\partial s_k^{\star}}{\partial p_j}$$

They will also satisfy absence of nominal illusion

$$s_i^{\star}(\mathbf{p}) = s_i^{\star}(\mathbf{p} + \delta)$$

 Given our assumptions on preferences, this result holds with correlation in unobserved tastes across products and in the mixed logit model More

## Symmetry

 With consideration sets, symmetry is violated and we suffer from nominal illusion

$$\frac{\partial s_j}{\partial p_k} \neq \frac{\partial s_k}{\partial p_j} 
s_j^*(\mathbf{p}) \neq s_j^*(\mathbf{p} + \delta)$$

- Changes in characteristics do not just impact utility, but also the probability of paying attention to particular subsets of goods
- We can use these asymmetries to identify attention probabilities,  $\mu(p_0)$  and  $\phi_j(p_j)$

## **Proof: Special Case**

- For purposes of this presentation, will walk through the proof of a special case of our more general framework
- ▶ Default specific model:  $\phi_j = 1$  for all j
- Choice probabilities take the form

$$s_0 = (1 - \mu) + \mu s_0^*$$
  
$$s_j = \mu s_j^*$$

where  $s_{i}^{\star} \equiv s_{i}^{\star} (\mathbf{p} | \{0,...,J\}).$ 

Changes in the characteristics of the default have two impacts on non-default goods:

$$\frac{\partial \mathbf{s}_j}{\partial \mathbf{p}_0} = \mu \frac{\partial \mathbf{s}_j^*}{\partial \mathbf{p}_0}$$

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$$\frac{\partial s_j}{\partial p_0} = \mu \frac{\partial s_j^*}{\partial p_0} + s_j^* \frac{\partial \mu}{\partial p_0}$$

$$\frac{\partial s_{j}}{\partial p_{0}} - \frac{\partial s_{0}}{\partial p_{j}} = s_{j}^{*} \frac{\partial \mu}{\partial p_{0}} + \mu \left( \frac{\partial s_{j}^{*}}{\partial p_{0}} - \frac{\partial s_{0}^{*}}{\partial p_{j}} \right)$$

$$= s_{j}^{*} \frac{\partial \mu}{\partial p_{0}}$$

$$= s_{j}^{*} \frac{\partial \mu}{\partial p_{0}} \frac{\mu}{\mu}$$

$$= \frac{s_{j}^{*}}{\mu} \frac{\partial \log(\mu)}{\partial p_{0}}$$

$$= s_{j} \frac{\partial \log(\mu)}{\partial p_{0}}$$

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$$\frac{\partial s_{j}}{\partial \rho_{0}} - \frac{\partial s_{0}}{\partial \rho_{j}} = s_{j}^{\star} \frac{\partial \mu}{\partial \rho_{0}} + \mu \left( \frac{\partial s_{j}^{\star}}{\partial \rho_{0}} - \frac{\partial s_{0}^{\star}}{\partial \rho_{j}} \right) \\
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Changes in consideration probabilities constructively identified by cross-derivative differences:

$$\frac{\partial \log(\mu)}{\partial p_0} = \frac{1}{s_j} \left[ \frac{\partial s_j}{\partial p_0} - \frac{\partial s_0}{\partial p_j} \right]$$

 Get the level of attention by integrating over the support of characteristics and pinning down the constant at point of symmetry

$$\mu = \exp\left(-\int \frac{1}{s_j} \left[\frac{\partial s_j}{\partial p_0} - \frac{\partial s_0}{\partial p_j}\right] dp_0\right)$$

Choice probabilities take the form

$$\begin{aligned} s_0 &= (1 - \mu) + \mu \sum_{C \in \mathbb{P}(0)} \prod_{l \in C} \phi_l \prod_{l' \notin C} (1 - \phi_{l'}) \, s_0^*(C) \\ s_j &= \mu \sum_{C \in \mathbb{P}(j)} \prod_{l \in C} \phi_l \prod_{l' \notin C} (1 - \phi_{l'}) \, s_j^*(C) \end{aligned}$$

for 
$$j > 0$$

- Need further source of variation in this model, with slight abuse of notation:
  - ▶  $s_j(\mathcal{J}/j')$ : market share of j when j' not available
  - NB Similar to Kawaguchi et al (MS, 2016) but without additional exclusion restriction
  - Can also express in terms of a full support assumption required for nonparametric identification of RUM

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Changes in consideration probabilities are the unique solution to a system of linear equations:

$$\frac{\partial s_j}{\partial \rho_0} - \frac{\partial s_0}{\partial \rho_j} = \frac{\partial \log(\mu)}{\partial \rho_j} s_j + \frac{\partial \log(\phi_j)}{\partial \rho_j} (s_0(\mathcal{J}/j) - s_0)$$
 where  $s_i = s_j(\mathbf{x}|\mathcal{J})$ .

► Final piece of puzzle: use nominal illusion to identify latent market shares

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► Final piece of puzzle: use nominal illusion to identify latent market shares

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#### **Estimation**

- Identification results constructive and so, in theory, consistent nonparametric estimators could be based on them
- However, in practice, nonparametric estimation is infeasible given the dimensionality problems
- Place a set of functional form assumptions on utility and the process driving consideration that are consistent with our framework
- We estimate special cases of our general framework in two scenarios, showing that asymmetries important for driving the ultimate results

## **Estimation: Application Assumptions**

Functional form assumptions a simple version of those followed in marketing literature and Goeree (2008):

$$s_{j}^{\star}(C) = \frac{exp(\alpha_{j} + x_{j}\beta)}{\sum_{j' \in C} exp(\alpha_{j'} + x_{j'}\beta)}$$
$$\phi_{j} = \frac{exp(\delta_{j} + x_{j}\gamma)}{1 + exp(\delta_{j} + x_{j}\gamma)}$$
$$\mu = \frac{exp(\delta_{0} + x_{0}\omega)}{1 + exp(\delta_{0} + x_{0}\omega)}$$

 Typical to estimate the parameters of the parametric model by maximum (simulated) likelihood (e.g. Goeree 2008)

#### Indirect Inference

- We instead pursue an estimation strategy that is grounded in the identifying variation at the heart of our identification proof
- Estimate the model by indirect inference
  - Match the parameters of a flexible auxiliary model that is able to capture cross-derivative asymmetries in the data
  - Intuitively, if estimate the auxiliary model on data simulated from the 'true' DGP, should get the same parameters as when estimating the auxiliary model on simulated data

## **Applications**

#### 1. Lab: Choice Experiment

- Alternative to a simulation exercise: we know that the model is misspecified
- We set the process generating which subset of 10 goods a respondent considers in a choice experiment
- Can we recover the parameters of this process without using information on what options a respondent considered?

#### 2. Field: Medicare Part D Choice

- Recent set of papers looking to disentangle switching costs and inattention in insurance choices
- Are the exclusion restrictions employed valid?
- Evaluation on an "active default" policy

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- Evaluation on an "active default" policy

### **Choice Experiment**

- Endowed respondents with \$25 and asked them to select their most preferred option from a set of goods that appeared on their screen
- 10 goods in full choice set chosen from Yale Bookstore with the price randomly drawn
- We set the probability that a particular good showed up on a respondent i's screen in round r as:

$$\phi_{j}(\rho_{ijr}) = \frac{exp\left(\delta_{j} + \gamma \rho_{ijr}\right)}{1 + exp\left(\delta_{j} + \gamma \rho_{ijr}\right)}$$

▶ Can we recover the (known)  $\delta_i$  and  $\gamma$ ?

### **Choice Experiment**



Collegiate Pacific Banner ("Yale University Lux et Veritas") \$8.00



Embroidered Towel From Team Golf \$20.00



Mug w/ Thumb Piece \$11.00



LXG Power Bank \$12.00



Moleskin Large Notebook with Debossed Wordmark, Unruled \$23.00

(You must wait 10 seconds before clicking next to make sure you consider all options)

Next

## **Choice Experiment**

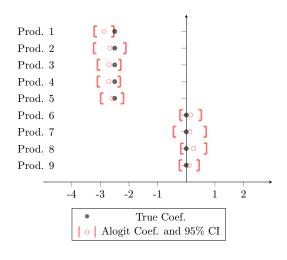
Auxiliary model specified as a flexible logit with good specific parameters:

$$\begin{aligned} v_{ijr} &= \omega_j + \theta_p p_{ijr} + \sum_{j'=0}^J \theta_{jj'} p_{ijr} p_{ij'r} \\ \widetilde{s}_{ijr} &= \frac{\exp\left(v_{ijr}\right)}{\sum_{j'} \exp\left(v_{ij'r}\right)} \end{aligned}$$

• Estimator of structural utility and consideration parameters,  $\psi = [\delta, \gamma, \alpha, \beta]$ , defined as:

$$\widehat{\psi} = \arg\min_{\psi} \left( \widehat{\theta^t} - \widehat{\theta^s}(\psi) \right)' W \left( \widehat{\theta^t} - \widehat{\theta^s}(\psi) \right)$$

#### Results: Attention Fixed Effects

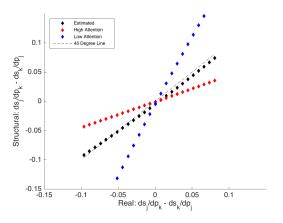


### Results: Price Coefficients

Table: Price Coefficients

	Conditional Logit	ALogit MLE	ALogit II	'Truth'
Utility	-0.054***	-0.1644***	-0.1284**	-0.173***
	(0.003)	(0.037)	(0.048)	(0.004)
Attention		0.137***	0.141***	0.15
		(0.017)	(0.025)	

# Asymmetries



### Field Application: Health Insurance

- Apply the Default Specific Model to Medicare Part D data on:
  - 20% sample of Part D beneficiaries from 2008-2009
  - Low Income Subsidy (LIS) beneficiaries "with stakes"
- DSC model applied by Heiss et al (2017) and Ho, Hogan & Scott-Morton although both rely on additional exclusion restrictions for identification
- Key question: how to explain inertia in choices over time?
- Often get implausibly large estimates of switching costs (> \$1,000)

#### Health Insurance

- Important for welfare evaluation of a smart default policy
  - Low switching because of high inattention?
  - Low switching because of utility relevant switching costs?
- Two sources of switching costs:
  - Paperwork costs, ρ: hassle and time to enrol in new scheme
  - Acclimation costs, α: cost of rescheduling deliveries and switching to new drugs
- Identification strategy:
  - Asymmetries: disentangle inattention from switching costs
  - $\blacktriangleright$  Random reassignment of LIS beneficiaries: separately identify  $\rho$  and  $\alpha$

## Utility

Choice probabilities in the DSC model given by:

$$m{s}_{\textit{ijt}} \equiv m{s}_{\textit{jt}}(m{x}_{\textit{it}}) = (\mathbf{1} - \mu_t(m{x}_{\textit{idt}})) \, \textit{Default}_{\textit{ijt}} + \mu_t(m{x}_{\textit{idt}}) m{s}_{\textit{it}}^\star(m{x}_{\textit{it}})$$

► Conditional on being awake, the utility of individual *i* from choosing plan *j* at time *t* is given by:

$$u_{ijt} = \mathbf{x}_{ijt}\beta + (\alpha + \rho)Default_{ijt} + \epsilon_{ijt}$$

When LIS beneficiaries no longer qualify for full premium subsidies, utility is given by:

$$u_{ijt} = x_{ijt}\beta + (\alpha + \rho)Default_{ijt} + \alpha \left(Default_{ij,t-1} \times Reassigned_{ijt}\right) + \epsilon_{ijt}$$

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#### Welfare

- Smart Default Policy: individuals are reassigned to an alternative plan and given the option of immediately switching back if they desire without enrolling in the new plan for a year.
- Assume that normative utility depends only on total cost and other observable factors
- Change in welfare associated with the policy can be expressed as:

$$\triangle W_i = W_i^1 - W_i^0$$

$$= \rho \left( s_{id}^1 - s_{io}^0 \right) + \alpha \triangle s_{io} + \sum_i \triangle s_{ij} v_{ij}$$
 (5.1)

### Preference Parameters: Medicare Part D

	Condition	al Logit	DSC N	Model
Utility:				
Annual Premium (hundreds)	-0.505***	(0.005)	-1.034***	(0.010)
Annual Out of Pocket Costs (hundreds)	-0.214***	(0.007)	-0.297***	(0.012)
Variance of Costs (millions)	2.246***	(0.089)	2.579***	(0.165)
Deductible (hundreds)	-0.516***	(0.009)	-0.724***	(0.013)
Donut Hole Coverage	0.691***	(0.027)	0.335***	(0.051)
Average Consumer Cost Sharing %	-1.181***	0.107	-4.128***	0.163
# of Top 100 Drugs in Formulary	0.038***	(0.004)	0.172***	(0.006)
Normalized Quality Rating	0.438***	(0.010)	0.515***	(0.015)
Original Plan	0.988***	(0.238)	1.314***	(0.257)
Assigned Plan	6.428***	(0.012)	4.240***	(0.078)
Acclimation Costs	\$196		\$127	
Paperwork Costs Attention Probability	\$1078		\$283 19.7%	

### Attention Parameters: Medicare Part D

	Conditional Logit		DSC Model	
Attention:				
Annual Premium (hundreds)	-	-	0.062***	(0.014)
Annual Out of Pocket Costs (hundreds)	-	-	0.030*	(0.012)
Variance of Costs (millions)	-	-	-0.627***	(0.159)
Deductible (hundreds)	-	-	0.069***	(0.020)
Donut Hole Coverage	-	-	-0.761***	(0.052)
Average Consumer Cost Sharing %	-	-	-1.447***	(0.219)
# of Top 100 Drugs in Formulary	-	-	-0.002	(0.010)
Normalized Quality Rating	-	-	-0.511***	(0.019)

Acclimation Costs	\$196	\$127	
Paperwork Costs	\$1078	\$283	
Attention Probability		19.7%	

## Welfare Simulations: Smart Default Policy 1

	Attention Cost				
	\$0	\$50	\$100	\$200	\$300
Conditional Logit Parameters	\$31	\$31	\$31	\$31	\$31
DSC Parameters	\$177	\$177	\$177	\$177	\$177
Direct Effect on Attention Probability 25% 50% 75% 100%					

# Welfare Simulations: Smart Default Policy 1

	Attention Cost				
	\$0	\$50	\$100	\$200	\$300
Conditional Logit Parameters	\$31	\$31	\$31	\$31	\$31
DSC Parameters	\$177	\$177	\$177	\$177	\$177
Direct Effect on Attention Probability 25% 50% 75% 100%	\$172 \$144 \$112 \$77	\$170 \$129 \$85 \$37	\$168 \$115 \$58 -\$2	\$164 \$86 \$4 -\$81	\$160 \$57 -\$50 -\$161

# Welfare Simulations: Smart Default Policy 2

	Attention Cost				
	\$0	\$50	\$100	\$200	\$300
DSC Parameters	\$222	\$222	\$222	\$222	\$222
Direct Effect on Attention Probability					
25%	\$215	\$213	\$210	\$204	\$199
50%	\$184	\$168	\$153	\$122	\$91
75%	\$150	\$122	\$95	\$39	-\$17
100%	\$114	\$74	\$35	-\$45	-\$124

## Overview of Additional Analysis

- Reduced form evidence of asymmetries: differential sensitivity of switching to changes in the characteristics of the default and rival plans
- Overidentification tests: test whether the exclusion restrictions used in the literature are valid

#### Conclusion

- ► Show identification of a class of consideration set models that are likely to be useful to applied researchers
- Exploit violations in symmetry of cross derivatives
- Assumptions already made by researchers in specifications with full-consideration typically sufficient for identification with limited consideration
- Demonstrate model utility/tractability through applications to choice in a variety of different settings including welfare evaluation of Smart Default Policy



# Symmetry Proof: Example Nested Logit

► To see in a very simple case, consider the nested logit in which cross-price effects take the form:

$$\frac{\partial s_{jm}}{\partial p_{km}} = \begin{cases} \beta s_{km} \left( \frac{\sigma}{1-\sigma} s_{jm|g} + s_{jm} \right) & \text{if } j \text{ and } k \text{ in the same nest} \\ \beta s_{jm} s_{km} & \text{otherwise} \end{cases}$$

- $s_{jm|g}$ : gives the within-nest market share of good j
- σ: how different substitution patterns are within and across nests.
- Clear that these are symmetric in products in different nests, but what about those in the same nest?

## Symmetry Proof: Example Nested Logit

For products in the same nest we have:

$$\frac{\partial s_{jm}}{\partial p_{km}} - \frac{\partial s_{km}}{\partial p_{jm}} = \beta \frac{\sigma}{1 - \sigma} \left( s_{km} s_{jm|g} - s_{jm} s_{km|g} \right)$$

Given that

$$s_{jm}=s_{jm|g}g_m$$

where  $g_m$  is the probability of buying a good from nest g. We have:

$$\frac{\partial s_{jm}}{\partial p_{km}} - \frac{\partial s_{km}}{\partial p_{jm}} = \beta \frac{\sigma}{1 - \sigma} \left( s_{km|g} g_m s_{jm|g} - s_{jm|g} g_m s_{km|g} \right)$$
$$= 0$$

# Symmetry Proof: General

▶ With [] denoting exclusion, the probability that option *j* is chosen under full consideration is given by:

$$\begin{split} s_{jm}^{\star} &= Pr\left(v_{jm} + \epsilon_{ijm} = \max_{j'} v_{j'm} + \epsilon_{ij'm}\right) \\ &= \int \int_{-\infty}^{v_{jm} + e - v_{0m}} ... \int_{-\infty}^{v_{jm} + e - v_{Jm}} f(z_0, ..., e, ..., z_J) dz_J ... [dz_j] ... dz_0 de \end{split}$$

This allows for an arbitrary correlation structure in the random utility errors.

# Symmetry Proof: General

► Then:

$$\frac{\partial s_{jm}}{\partial p_{j'm}} = -\beta \int \int_{-\infty}^{v_{jm}+e-v_{0m}} ... \left[ \int_{-\infty}^{v_{jm}+e-v_{jm}} \right] ... \left[ \int_{-\infty}^{v_{jm}+e-v_{j'm}} \right] ... \int_{-\infty}^{v_{jm}+e-v_{Jm}} f(z_0, ..., e, ..., v_{jm} + e - v_{j'm}, ..., z_J) dz_J ... [dz'_j] ... [dz'_j] ... dz_0 de$$

# Symmetry Proof: General

▶ Using the change of variables  $t = v_{jm} + e - v_{j'm}$ , one obtains:

$$\frac{\partial s_{jm}}{\partial p_{j'm}} = -\beta \int \int_{-\infty}^{v_{j'm}+t-v_{0m}} .. \left[ \int_{-\infty}^{v_{j'm}+t-v_{j'm}} \right] .. \left[ \int_{-\infty}^{v_{j'm}+t-v_{jm}} \right] .. \int_{-\infty}^{v_{j'm}+t-v_{jm}} f(z_0, ..., v_{j'm} + t - v_{jm}, ..., t, ..., z_J) dz_J ... [dz'_j] ... [dz_j] ... dz_0 dt$$

$$= \frac{\partial s_{j'm}}{\partial p_{jm}}$$

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